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Abstract

The study examines the effects of accessibility to the nearest urban metropolitan area on rural poverty by using Japanese municipality-level data. We conduct nationwide cross-sectional analyses, and find that a larger time distance to the nearest urban metropolitan area significantly increases regional poverty rates. In addition, the study focuses on opening of new commuting train, Tsukuba Express (TX), connecting Tokyo and Ibaraki prefecture, a suburban area of Tokyo. We conduct municipality-level panel analyses, and the results suggest that opening TX reduced rural poverty rates of the surrounding areas, but the effects required 6–10 years to be observed. Therefore, regional policy makers might need to consider that transportation investments that improve inter-regional accessibility do not affect regional economic performance for several years.

JEL classification: R11, R12, R13, R41, and R42

1. Introduction

Even in developed countries, poverty remains a serious problem. Candy and Smith (2014) compare ten different definitions of absolute poverty rates for the United States and point out that one index of absolute poverty rates reaches roughly five percent.¹ In Japan, the government provides public assistance to people in poverty who cannot pay for minimum costs of living, and the share of people receiving public assistance increased from 0.70% in 1995 to 1.70% in 2015.² Given this we can conclude that there are still many poor people who cannot support themselves without assistance, and the number has tended to increase more recently. Therefore, we must investigate factors shaping poverty in developed countries to reduce poverty rates.

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¹ Although the absolute poverty rate is officially defined as the share of people living with an income of less than 1.90 dollars per day, the amount is less than 2.00 dollars per day according to Candy and Smith. The index of absolute poverty rate of roughly 5% corresponds the definition of absolute poverty given in Shaefer and Edin (2013). The index focuses only on households with children and excludes the effects of government assistance such as food stamps. As a source, the estimation adopts the Survey of Income and Program Participation.

² Roughly 2,140,000 people received public assistance in 2015 according to the Ministry of Health, Labor and Welfare.

The ILO (2016) announced that improving income levels is crucial to reducing poverty rates. As empirical evidence of the relationship between income levels and poverty, Förster and d'Ercole (2005) examine OECD countries for the second half of the 1990s and find that poverty rates and income levels are strongly correlated. From this discussion factors affecting regional income levels may also affect regional poverty levels.

As an important factor that affects regional economic performance, we consider agglomeration spillover effects. Marshall (1920) mentions the possibility that firms tend to concentrate to secure advantages such as rich labor markets, low transportation costs of inputs and outputs, and knowledge spillovers. Such benefits of agglomeration increase the productivity of firms in urban metropolitan areas, and the effects are known to spill over to surrounding regions.

The magnitude of agglomeration spillover effects is known to decrease with distance from urban metropolitan areas. Rosenthal and Strange (2003, 2006) empirically investigate six industries in the United States. They suggest that the amount of employment rapidly decreases with distance from agglomerations in five of six industries.

The above discussion suggests that regional economic performance may diminish with distance from urban metropolitan areas. In terms of poverty levels, Partridge and Rickman (2008) investigate the relationship between regional poverty rates and distance from the nearest urban metropolitan areas by using county-level data for the United States and show that a larger linear distance from a nearby urban metropolitan area increases poverty rates in counties. Partridge and Rickman interpret the heterogeneous distribution of poverty as a result of decreases in regional labor demand and wage levels with distance from economic agglomerations.

In this study we estimate the effects of accessibility to the nearest urban metropolitan area on regional poverty by using Japanese municipal-level data. From our nationwide cross-sectional analyses we find that a larger time distance to the nearest urban metropolitan area significantly increases regional poverty rates. The monetary magnitude is such that a one-minute increase of time distance to the nearest urban metropolitan area increased the number of households in poverty by roughly 0.78 and the annual expenditure of regional governments for public assistance by approximately 1.75 million yen on average in 2014.

In addition, we focus on the case of a new commuting train that opened in 2005, the Tsukuba Express (TX), which connects Tokyo and suburban areas, and we conduct panel analyses to understand the impacts of changing levels of accessibility to closely located urban metropolitan areas on rural poverty. From our panel analyses we find that improvements in accessibility to the nearest urban metropolitan area significantly decrease poverty rates and even when controlling for municipality fixed effects. We also

find that the effects of reducing regional poverty are observed in municipalities located close to TX. This result is consistent with our hypothesis that improvements in accessibility to the nearest urban metropolitan area will spread the range of positive spillover effects from urban metropolitan areas, stimulate their economic performance, and reduce poverty levels.

In a related study, Partridge and Rickman (2008) investigate the relationship between regional poverty rates and distance from the nearest urban metropolitan areas in the United States. They show that a larger linear distance from a nearby urban metropolitan area increases poverty rates in counties. However, two issues are not considered in their study. First, Partridge and Rickman use linear distance as a distance variable, which cannot measure accurate interregional accessibility.³ To solve this problem, we adopt time distance, which can measure actual transportation costs as an interregional accessibility variable, as well as linear distance. From our estimations, a larger time distance to the nearest urban metropolitan area increases regional poverty rates while linear distance to the nearest urban metropolitan area does not shape rural poverty. Second, Partridge and Rickman (2008) only conduct a cross-sectional analysis, and their results may contain biases resulting from neglecting unobservable regional characteristics. Against this background we conduct a panel analysis to understand the effects of changing transportation costs for traveling to closely located urban metropolitan areas on rural poverty rates while controlling for time invariant regional characteristics, and we find that the opening of a new commuting train reduces poverty levels in regions located close to the commuting train.

In another related study about the location of poverty, Glaeser, Kahn and Rappaport (2008) investigate the distribution of poverty in areas roughly 16 kilometers from a CBD in the United States, and they find that those living in poverty tend to live in central areas of cities and to enjoy the advantages of better public transportation infrastructure. They also point out that regional median income decreases with distance to a CBD. However, Clark, Huang and Withers (2003) find that more than one quarter of employees in the Seattle labor market had a commute distance of greater than 16 kilometers from their residences in the 1990s. Since commutable areas have expanded with public transportation and residential development specifically in developed countries, it may be that the range used by Glaeser, Kahn and Rappaport are not sufficient to consider the distribution of poverty in surrounding areas of urban metropolitan areas.

³ Boscoe et al. (2012) focus on the relationship between housing prices and distance to the nearest hospital in reference to the United States and Puerto Rico. They find that linear distance does not work appropriately as a measure of accessibility when there are geographic barriers that prevent people from traveling.

The structure of this paper is as follows. The next section describes mechanisms of the relationship whereby access to urban metropolitan areas affects rural poverty in reference to previous studies. Section 3 describes our estimation models and variables. Section 4 describes the results of our cross-sectional analysis. Section 5 describes the configuration of the panel analysis and its results. Section 6 concludes.

2. Background mechanism

This section describes the mechanism whereby regional accessibility to closely located urban metropolitan areas affects rural poverty referring to previous studies. The seriousness of poverty conditions in a region depends on the regional wage and employee level, which is determined as the equilibrium of regional labor demand and labor supply. Partridge and Rickman (2008) formulate a relationship whereby regional employment and wage rates affect regional poverty, which can be written as the following function.

$$\text{Poverty}_i = f_i^{pov}(er_i, wr_i, \text{other}_i^{pov}), \quad (1)$$

where er_i is the employment rate of region i and where wr_i is its wage rate. other_i^{pov} is the vector of other variables that affect the poverty conditions of region i . To understand how Poverty_i has an effect, we consider the effects of er_i and wr_i .

The employee and wage rates of a region depend on the interaction of regional labor demand and labor supply. These relationships are written as the following functions.

$$er_i = f_i^{er}(l_i^d, l_i^s), \quad (2)$$

$$wr_i = f_i^{wr}(l_i^d, l_i^s), \quad (3)$$

where l_i^d is the labor demand of region i , and where l_i^s is the labor supply. When other factors are given, an increase in l_i^d increases er_i or wr_i , and an increase in l_i^s decreases er_i or wr_i . Then, we consider factors that determine the level of regional labor demand and labor supply.

On the labor demand side, agglomeration economy spillovers and increases in labor demand of surrounding regions occur, and spillover effects are known to diminish with distance. Audretsch et al. (2005) focus on the location of high technology-based firms in the United States and find that high

technology-based firms heavily concentrate within 50 kilometers of universities. The above results suggest that the level of regional labor demand decreases with distance from economic agglomeration.

On the labor supply side, rural workers are known to experience difficulties in accessing urban labor markets across regions. Lucas (2001) introduces evidence suggesting that rural workers remain in their own areas in spite of higher income levels in urban metropolitan areas. Molho (1995) provides evidence referenced in Lucas (2001) suggesting that rural workers tend to remain in rural areas due to their attachments to the culture or human relations in areas in which they live. In addition, Lucas (2001) identifies costs of information about urban labor markets, which increase with distance from urban metropolitan areas, as a reason for why rural workers remain in the areas in which they live.

The above results suggest that while labor demand concentrates in urban metropolitan areas and while this demand rapidly diminishes with distance from economic agglomerations, rural worker mobility is low, and such workers experience difficulties in migrating to urban metropolitan areas across regions in pursuit of higher wages. This causes labor demand to diminish with distance from economic agglomerations with a larger slope than that of labor supply. We can express these relationships as the following functions:

$$l_i^d = f^{ld}(Distance_i^{UA}, other_i^{ld}), \quad \frac{\partial l_i^d}{\partial Distance_i^{UA}} < 0, \quad (4)$$

$$l_i^s = f^{ls}(Distance_i^{UA}, other_i^{ls}), \quad \frac{\partial l_i^s}{\partial Distance_i^{UA}} < 0, \quad (5)$$

$$\frac{\partial l_i^d}{\partial Distance_i^{UA}} < \frac{\partial l_i^s}{\partial Distance_i^{UA}}, \quad (6)$$

where $Distance_i^{UA}$ is the distance to the nearest urban metropolitan area of region i. From functions (2) and (3), increases in $Distance_i^{UA}$ decrease regional wages or employee levels through changes in l_i^d and l_i^s , and from function (1) increases in $Distance_i^{UA}$ worsen regional poverty conditions. The relationship can be written as the following function:

$$poverty_i = f_i^{pov}(Distance_i^{UA}, other_i), \quad \frac{\partial poverty_i}{\partial Distance_i^{UA}} > 0, \quad (7)$$

where $poverty_i$ is the poverty rate of region i.

From the above discussion we assume that stronger accessibility to closely located economic agglomerations improves regional poverty conditions. We examine the impact of distance to the nearest urban metropolitan area on rural poverty rates with Japanese municipality level data.

3. Cross-sectional analysis

3.1. Empirical strategy

Based on the above theoretical background, this section describes the estimation model that explains municipalities' poverty rates. The estimation model is as follows:

$$\text{Pov}_i = \alpha_i + \beta \text{Distance}_i + \delta \mathbf{X}_i + \theta \text{Prefecture}_i + u_i. \quad (8)$$

Pov_i is municipality i 's poverty rate. Distance_i is municipality i 's distance to the nearest urban metropolitan area. According to the above discussion, we expect that β to be negative.

\mathbf{X}_i is the vector with variables relating to municipality i 's poverty rate. It includes three types of variables explaining municipality i 's population structure, economic activity, and education level.

Prefecture_i is a prefecture dummy in which municipality i is contained. We use it to control for heterogeneity comes from the prefecture of municipality i . u_i is the error term.

3.2. Data

This section describes the data that we use. First, we clarify the definition of municipalities and urban metropolitan areas.

The Japanese government defines a *municipality* as the smallest unit of an administrative district composed of cities, towns, villages, and specified districts.⁴ We define urban metropolitan areas as municipalities with populations of over 300,000; this is a condition of the *core city*, which is a legal urban metropolitan area determined by article 252 of the Local Autonomy Law.⁵ In Japan, there were 71 urban metropolitan areas in 2012. In addition, we regard 23 specified districts in the Tokyo metropolitan area as one urban metropolitan area. In this study, each municipality has a nearest urban metropolitan area. The nearest urban metropolitan area of a municipality is defined as the urban metropolitan area at the closest linear distance to a rural municipality.

Since we cannot observe the distribution of poverty in each municipality and the distance between each household in urban areas and the center of an urban metropolitan area, we exempt municipalities that are urban metropolitan areas from our sample. In this study, we focus on the distribution of poverty in suburban and rural areas.

⁴ In 2012, there were 1,747 municipalities that included 786 cities, 754 towns, 184 villages, and 23 specified districts in the Tokyo metropolitan area.

⁵ In 2016, the definition of a core city was changed to a city with a population of over 200,000. However, we adopt the previous definition used in 2012, the period of our sample.

Distance_{*i*} is municipality *i*'s distance to the nearest urban metropolitan area. In our cross-sectional analysis, we adopt linear distance, time distance by car, and time distance by public transport as accessibility variables to the nearest urban metropolitan area.

Linear distance is measured as the linear distance in kilometers between the government offices of a rural municipality and the municipality's nearest urban metropolitan area. We calculate linear distances between municipalities using location-based and coordinate conversion services provided by the Geospatial Information Authority of Japan.

Time distance by car measures how long people must spend to travel between two government offices by car. This variable is calculated using Google Maps.⁶ Time distance by public transport is the amount of time people must spend to travel between the municipality in which they live and its nearest urban metropolitan area by train and bus.⁷ We calculate this variable with Timetables published by the Japan Travel Bureau (JTB), a representative timetable of public transportation in Japan. For municipalities that include stations, we measure the time distance between their representative station, which is defined in Timetables, and the representative station of its nearest urban metropolitan area.⁸ For municipalities without stations we add the time distance between their government offices' nearest bus stops and the nearest stations to the time distance for traveling from a rural station to the representative station of the nearest urban metropolitan area. We calculate the optimal path between a rural municipality and its nearest urban metropolitan area for commuters.⁹ The unit of time distance is one minute.

As a measure of regional poverty rates we adopt a municipality's share of households receiving public assistance. This is defined as the ratio of households receiving public assistance to 100 households in each municipality. To receive public assistance a household must live under the poverty line, which is determined by standards created by the Ministry of Health, Labour and Welfare (MHLW). Each municipality's income threshold for providing public assistance controls for each municipality's price level determined by the MHLW. We regard municipalities' shares of households receiving public assistance work as a proxy for the absolute poverty rate, which is defined as "*the inability to meet basic*

⁶ We obtain data for 2016.

⁷ Each municipality's principal station is defined by its government.

⁸ In a cross-sectional analysis we refer Timetables (2010) records for time distance by public transport for April 1st, 2010.

⁹ The optimal path is calculated as the fastest path from a rural station to the representative station of the nearest urban metropolitan area, but we exempt limited express trains and bullet trains, which take higher fares as commuting methods.

needs of health and nutrition” (Deaton, 2004, pp11). As an additional reason to adopt public assistance, little municipal-level data on poverty in Japan are available. To calculate the public assistance rate of each municipality, we use Prefectural Statistic Manuals (2012).¹⁰

Table 2.1 shows the number and share of reasons for discontinuing public assistance to households in Japan as a whole between 2012 and 2016. For all types of households, the share of public assistance removed due to increasing employment income is not a considerable at between 13.9 and 16.0%. The main process affecting this category is the death of a receiver (28.6-34.0%) and mostly for elderly households and households with handicapped members not in the labor force. Without these effects for fatherless households and other households, the main cause of discontinued public assistance is an increased income or income from a job (29.5-36.0% for other households). Given this data, we note that many households receive public assistance due to receiving little or no income even when active in the labor force, and that improvements in municipal wage or employee levels could reduce the share of households receiving public assistance.

¹⁰ The Prefectural Statistic Manuals (2012) record each dataset for 2011.

Table 1. The Number and Share of Households Discontinuing PA in Japan

	2012	2013	2014	2015	2016
<i>All Types of Households</i>					
All	13,986	11,901	11,710	13,333	11,474
Death	4,002 (28.6)	3,639 (30.6)	3,670 (31.3)	4,342 (32.6)	3,900 (34.0)
Increasing or Deriving Income by Jobs	1,974 (14.1)	1,659 13.9	1,878 16.0	2,127 16.0	1,826 15.9
<i>Elderly Households</i>					
All	4,895	4,430	4,524	5,550	5,104
Death	2,810 (57.4)	2,600 (58.7)	2,752 (60.8)	3,398 (61.2)	3,110 (60.9)
Increasing or Deriving Income by Jobs	85 (1.7)	83 (1.9)	102 (2.3)	110 (2.0)	122 (2.4)
<i>Fatherless Households</i>					
All	972	717	719	888	665
Death	40 (4.1)	7 (1.0)	5 (0.7)	7 (0.8)	7 (1.1)
Increasing or Deriving Income by Jobs	224 (23.0)	164 (22.9)	189 (26.3)	271 (30.5)	214 (32.2)
<i>Households with Handicapped Members</i>					
All	4,061	3,168	2,858	3,236	2,648
Death	933 (23.0)	823 (26.0)	744 (26.0)	760 (23.5)	636 (24.0)
Increasing or Deriving Income by Jobs	368 (9.1)	353 (11.1)	361 (12.6)	417 (12.9)	380 (14.4)
<i>Other Households</i>					
All	4,058	3,586	3,609	3,659	3,057
Death	219 (5.4)	209 (5.8)	169 (4.7)	177 (4.8)	147 (4.8)
Increasing or Deriving Income by Jobs	1,297 (32.0)	1,059 (29.5)	1,226 (34.0)	1,329 (36.3)	1,110 (36.3)

Percentages of the reasons are in parentheses.

Partridge and Rickman (2008) use county-level poverty rates based on the poverty standard defined by the U.S. Census Bureau. The standard aims to identify whether a household is in absolute poverty, and it controls for states' price levels and for the number of household members; it applies similar requirements to those of the Japanese income threshold for receiving public assistance. Similarities between the two poverty standards allow us to easily compare our estimation results to those of Partridge and Rickman (2008).

Using \mathbf{X}_i we control for factors relating to municipality i 's poverty rate. It covers three types of variables. (I) Variables that explain municipality i 's population structure include the number of households and age structures (the share of the population under 15 and the share of the population over 65). (II) Variables on municipality i 's economic performance include industrial structure (the share of laborers in the primary and manufacturing sectors) and municipality i 's unemployment rate.

(III) Variables reflecting municipalities' education levels include the share of people who have graduated from a university and the share of high school graduates. To obtain these variables, we use the Statistical Observations of Prefectures provided by the Ministry of Internal Affairs and Communications Statistics Bureau (2010).¹¹

Table 2.2 presents the summary statistics. We use 33 prefectures from a total of 47 reporting the share of households receiving public assistance at the municipal level.¹² We find from the table that the mean municipal linear distance to the nearest urban metropolitan area is roughly 35 kilometers, and time distance is roughly 50 minutes. This indicates that we can focus on the distribution of poverty in suburban areas unlike the geographical range examined by Glaeser, Kahn and Rappaport (2008), we investigate distributions of the poor in urban metropolitan areas. The table also shows that the mean poverty rate is approximately 2.3%. This is much lower than the relative poverty rate in Japan of roughly 16% for 2012 announced by the OECD. We thus focus only on households in serious poverty which is difficult to live under with minimum standards without receiving public assistance.

¹¹ The Statistical Observations of Prefectures (2010) records data for 2010.

¹² The 33 available prefectures include Aichi, Chiba, Fukui, Fukuoka, Fukushima, Gifu, Hiroshima, Hokkaido, Hyogo, Ibaraki, Iwate, Kagawa, Kagoshima, Kanagawa, Kochi, Kumamoto, Kyoto, Miyazaki, Nagasaki, Nara, Oita, Okinawa, Osaka, Saga, Saitama, Shiga, Shimane, Tochigi, Tokyo, Tottori, Toyama, Wakayama, and Yamaguchi.

Table 2. Summary Statistics

	Mean	Std. Dev.	Min	Max
<i>DISTANCE to the Nearest Urban Metropolitan Area</i>				
Linear Distance (km)	35.6	39.8	4	412
Time Distance by Car (min)	50.3	37.2	9	338
Time Distance by Public Transport (min)	52.7	53.0	3	510
<i>Poverty</i>				
Household Receiving PA	873	1,008	35	9,384
Poverty Rates (%)	2.3	1.5	0	9
<i>Other Regional Characteristics</i>				
Population	92,182	61,673	4,387	290,959
Household	36,664	25,682	2,000	126,180
Share of Population Under 15	13.1	1.9	6	20
Share of Population Over 65	26.2	5.2	14	46
Share of Workers in Primary Sector	26.6	29.5	0	80
Share of Workers in Industry Sector	25.5	7.1	11	49
Share of Unemployed	3.4	0.8	2	7
Share of High School Graduates	36.0	7.1	17	97
Share of University Graduates	11.9	5.0	3	33
N = 378				

4. Cross-sectional results

This section describes the results of our cross-sectional estimation. Table 2.3 shows the estimation results. Column (1) uses linear distance as Distance_i , and its coefficient is positive and insignificant. This result suggests that linear distance to the nearest urban metropolitan area does not have significant effects on rural poverty rates. Column (2) adopts time distance by car as Distance_i . The coefficient is positive with 10% statistical significance. For magnitude we find that one-minute increases of time distance by car to the nearest urban metropolitan area increase rural poverty rates by roughly 0.003 percentage points. Column (3) shows the estimation of time distance by public transport, and its coefficient is positive with 5% statistical significance. This result is roughly the same as that of estimation (2). One-minute increases of time distance by public transport to the nearest urban metropolitan area increase rural poverty rates by roughly 0.002 percentage points.

Columns (1)–(3) are fundamentally consistent with the results of Partridge and Rickman (2008) in suggesting that a longer distance to the nearest urban metropolitan area will worsen rural poverty conditions. However, our results also show that there is a difference between the significance of linear

distance and time distance. From columns (1)–(3) we find that while a municipality’s linear distance to the nearest urban metropolitan area does not have significant effects on rural poverty rates, time distance significantly affects rural poverty rates. These results suggest that while linear distance does not precisely capture interregional accessibility in regions such as Japan that have many geographical barriers, time distance can measure accessibility.

We consider the monetary impact of distance to the nearest urban metropolitan area on rural governments. From column (3) the magnitude of a one-minute increase of time distance represent roughly 0.002 percentage point increase in rural poverty rates. These results show that, for instance, an increase of 10 minutes of time distance to the nearest urban metropolitan area causes municipalities’ annual expenditures on public assistance to increase by approximately 17.5 million yen on average.¹³

¹³ In 2014 there were roughly 56.4 million households in Japan and roughly 1.6 million households received public assistance; thus, the share of households receiving public assistance was approximately 2.38%. There were 1,718 municipalities in 2014, and each municipality includes roughly 931 households receiving public assistance on average. Total expenditures of the Japanese government on public assistance amounted to roughly 3,843 billion yen, and when dividing total expenditures by the number of municipalities, a municipality’s average expenditures on public assistance was roughly 2.2 billion yen. From the average number of households receiving public assistance and from the size of expenditures, the average level of public assistance dedicated to each household amounted to roughly 2.4 million yen. From the results of (2) and (3), an increase of 10 minutes of time distance increases the share of public assistance by roughly 0.02; this adds roughly 7.82 households with public assistance. From this result, a municipality’s annual expenditures for public assistance increases by roughly 17.5 million yen on average.

Table 3. Results of the Cross-Sectional Analysis

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
	<i>Baseline Sample</i>			<i>Urban Metropolitan Areas with Populations Larger than 500,000</i>			<i>Urban Metropolitan Areas with Populations Smaller than 500,000</i>		
ln (Linear Distance)	0.078 (0.071)	–	–	0.23 (0.22)	–	–	0.066 (0.074)	–	–
Time Distance by Car	–	0.0032* (0.0012)	–	–	0.0020 (0.0012)	–	–	0.0018* (0.00075)	–
Time Distance by Public Transport	–	–	0.0018** (0.00086)	–	–	0.0015 (0.00085)	–	–	0.0018* (0.00075)
Population (thousand)	0.0020*** (0.00059)	0.0019** (0.00055)	0.0021*** (0.00058)	0.00079 (0.00060)	0.00079 (0.00062)	0.00088 (0.00060)	0.0011** (0.00038)	0.0011** (0.00038)	0.0011** (0.00038)
Share of Population Under 15	-0.062 (0.039)	-0.051 (0.038)	-0.06 (0.038)	-0.082 (0.053)	-0.068 (0.054)	-0.072 (0.052)	-0.048 (0.034)	-0.038 (0.034)	-0.039 (0.034)
Share of Population Over 65	-0.017 (0.018)	-0.010 (0.018)	-0.016 (0.017)	-0.067** (0.025)	-0.060* (0.024)	-0.064** (0.024)	-0.022 (0.015)	-0.019 (0.015)	-0.020 (0.015)
Share of Workers in Primary Sector	-0.0029 (0.0017)	-0.0043* (0.0017)	-0.0036* (0.0018)	-0.0016 (0.0018)	-0.0032 (0.0019)	-0.0028 (0.0019)	-0.0030** (0.0011)	-0.0026* (0.0011)	-0.0026* (0.0011)
Share of Workers in Industry Sector	-0.0040 (0.0070)	-0.0058 (0.0066)	-0.0039 (0.0067)	-0.017* (0.0080)	-0.019* (0.0079)	-0.017* (0.0078)	-0.012* (0.0052)	-0.010* (0.0051)	-0.010* (0.0051)
Unemployed Rates	0.15* (0.062)	0.15* (0.062)	0.14* (0.062)	-0.016 (0.060)	-0.023 (0.053)	-0.029 (0.052)	0.083* (0.038)	0.075* (0.036)	0.075* (0.036)
10 Years' Lag of Poverty Rates	0.80*** (0.076)	0.78*** (0.075)	0.79*** (0.075)	1.34*** (0.10)	1.36*** (0.10)	1.34*** (0.10)	1.08*** (0.053)	1.11*** (0.047)	1.11*** (0.047)
Share of High School Graduates	0.0037 (0.0082)	0.0057 (0.0080)	0.0052 (0.008)	0.019* (0.0087)	0.023* (0.0093)	0.023* (0.0092)	-0.0018 (0.0065)	0.00049 (0.0062)	0.00049 (0.0063)
Share of University Graduates	-0.010 (0.012)	-0.0074 (0.012)	-0.0078 (0.012)	-0.011 (0.018)	-0.019 (0.014)	-0.017 (0.014)	-0.027** (0.0093)	-0.023* (0.0092)	-0.023* (0.0092)
Intercept	1.33 (1.05)	0.93 (1.06)	1.25 (1.05)	2.77 (1.61)	2.75 (1.47)	2.83 (1.45)	1.92 (1.00)	1.57 (0.98)	1.57 (0.98)
Prefecture	Y	Y	Y	Y	Y	Y	Y	Y	Y
Adjusted R ²	0.76	0.77	0.77	0.92	0.92	0.92	0.93	0.93	0.94
N	372	368	372	171	170	171	197	194	194

Signif. codes: 0 '***' 0.01 '**' 0.05 '*' 0.1. We adopt a cluster robust standard error.

Cluster robust standard errors are in parentheses.

It may be that the scale of economic agglomeration affects the magnitude of spillover effects on surrounding regions. In analyses (4)–(9) we divide municipalities by the scale of the population of their nearest urban metropolitan areas. We define the threshold of municipalities as whether their nearest urban metropolitan area has a population of over 500,000; this is one of the requirements for designation as an ordinance designated city, a representative system for define metropolises in Japan. Columns (4)–(6) present the results of estimations using municipalities whose nearest urban metropolitan area has a population of larger than 500,000. Columns (7)–(9) cover use municipalities with populations of less than 500,000.

As a distance variable, linear distance is used in column (4), time distance by car is adopted in column (5), and time distance by public transport is used in column (6). The results of columns (4)–(6) suggest that distance to the nearest urban metropolitan area does not have significant effects on rural poverty rates in municipalities whose nearest urban metropolitan area has a population larger than 500,000. From

columns (7)–(9) we observe qualitatively similar results to those of columns (1)–(3) in that the coefficient of linear distance is positive and insignificant, and the coefficients of time distance by car and public transport are positive and statistically significant. Column (7) uses linear distance as Distance_i , and its coefficient is positive and insignificant. Column (8) adopts time distance by car as Distance_i . The coefficient is positive with 10% statistical significance. This result shows that one-minute increases of Distance_i increase rural poverty rates by roughly 0.002 percentage points. Column (9) provides the estimation with time distance by public transport, and its coefficient is positive with 10% statistical significance. This result suggests that one-minute increases of Distance_i cause rural poverty rates to increase by roughly 0.002 percentage points. Columns (4)–(9) show that whereas distance to the smaller urban metropolitan areas impacts rural poverty rates, accessibility to larger urban metropolitan areas does not have significant effects.

We consider why only municipalities closely located to a smaller urban metropolitan area experience significant effects of distance to urban metropolitan areas. From columns (5) and (8) we find that in municipalities close to a larger urban metropolitan area, the magnitude of the coefficient (roughly 0.002) is larger than that of municipalities close to a smaller urban metropolitan area (roughly 0.0018). However, the coefficient for municipalities close to a larger urban metropolitan area has a considerably higher standard error (roughly 0.0012) than that of municipalities close to a smaller urban metropolitan area (roughly 0.00075), resulting in the insignificance of time distance shown in column (5). From columns (5), (6), (8), and (9) we observe that municipalities close to larger urban metropolitan areas consistently have a larger standard error of time distance to the nearest urban metropolitan areas than municipalities close to smaller urban metropolitan areas. This suggests that there are broader heterogeneities in the impacts of agglomeration economies from the nearest urban metropolitan area among municipalities positioned close to larger urban metropolitan areas than among those positioned close to smaller urban metropolitan areas; for example, only some of the municipalities receive significant benefits through the strong accessibility to the nearest urban metropolitan area while the others do not. We consider the possibility that such heterogeneities are caused by urban shadow effects whereby an urban metropolitan area reabsorbs economic activities of surrounding regions, resulting in their economic declines. Larger metropolitan areas may present more serious urban shadow effects on their surrounding areas, resulting in insignificant effects of accessibility to the nearest urban metropolitan area with more than 500,000 residents on poverty levels in surrounding areas.

5. Panel analysis

Although our estimation model includes all possible variables to control for municipality heterogeneity, our model may still miss unobservable characteristics that cannot be controlled by these variables, and the results may include some biases. To control for unobservable time invariant heterogeneities of municipalities, we conduct a panel analysis and estimate the impact of improving accessibility to the nearest urban metropolitan area on rural poverty rates.

5.1. Opening of TX

To control for unobservable and time invariant characteristics of municipalities, we focus on the opening of a new commuting train running between Tokyo and its surrounding cities, Tsukuba Express (TX), in 2005. TX connects the city of Tsukuba in Ibaraki prefecture, a northern suburban area of Tokyo, to the Akihabara area located in the Tokyo CBD. The opening of TX shortened the time distance between Tsukuba and Akihabara from 60 minutes to 40 minutes. This drastic change in accessibility to urban metropolitan areas might affect the geographic range of spillover effects of economic agglomeration on regions surrounding the rail line. For instance, households' job selection behaviors or firms' location decisions may reflect the impacts of changing levels of accessibility in municipalities surrounding the railway. Since the railway also includes a station in Kashiwa (an urban metropolitan area in Chiba prefecture), the opening of TX might have impacts resulting from a change in accessibility to other smaller urban metropolitan areas and to Tokyo. Figure 2.1 shows routes of TX and the Joban line, an existing railway connecting municipalities in Ibaraki to Tokyo. The TX route is shown in red, and the Joban line is shown in blue. Each circle represents a railways station.

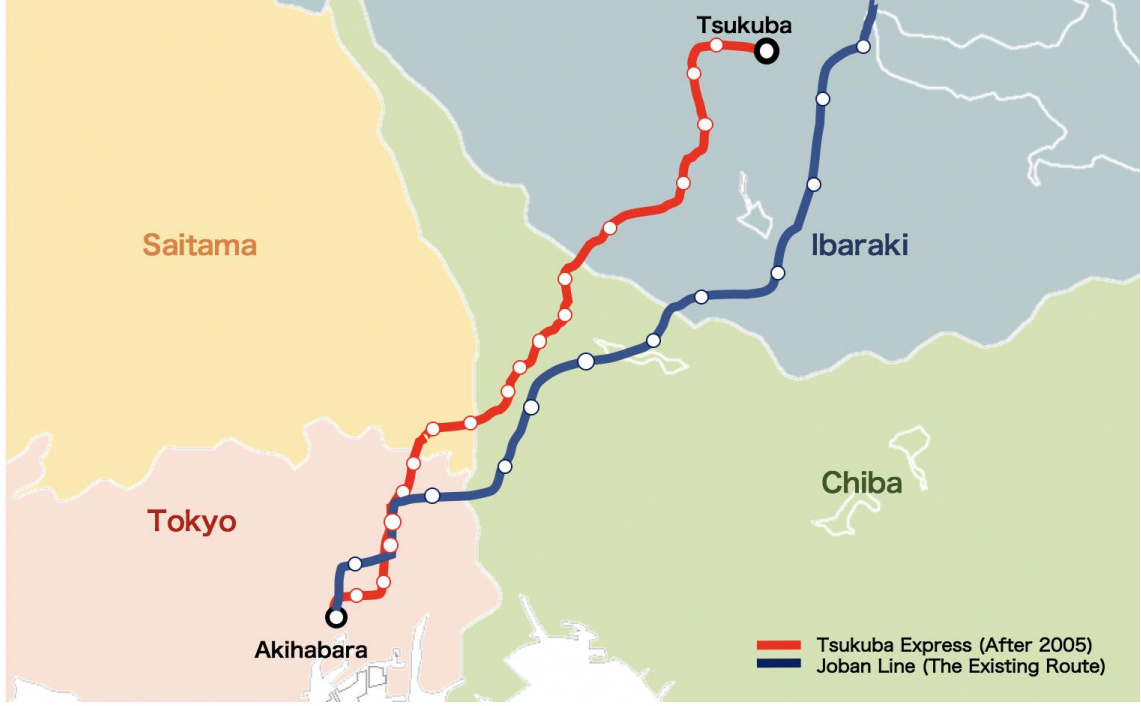


Figure 1: Routes of the Tsukuba Express and Joban Line.

5.2. Panel analysis

We create panel data and estimate the impacts of accessibility to urban metropolitan areas on municipalities' poverty rates while controlling for the time invariant characteristics of municipalities. To create the panel data, we use JTB's Timetables (JTB, 2000, 2005, 2010, 2015) and calculate the time distance between stations using the appropriate commuting route. The estimation model is as follows:

$$\text{Pov}_{it} = \alpha_{it} + \beta \text{TimeDistance}_{it} + \delta \mathbf{X}_{it} + \eta_i + \varepsilon_{it}. \quad (9)$$

η_i denotes municipality fixed effects. ε_{it} is the error term derived from the time variant characteristics of municipality i .



Figure 2. Regions of the Ibaraki Prefecture and the Route of the Tsukuba Express

In our panel analysis only time distance by public transport is used as distance variable, as linear distance cannot identify changes in interregional accessibility. Additionally, we cannot observe the previous time distance by car.

Ibaraki prefecture is divided into five regions: the central, southern, western, northern, and Rokko regions.¹⁴ TX terminates at the city of Tsukuba in the southern region. Figure 2.2 shows the main suburban cities of Mito, Hitachi and Tsukuba: five regions of Ibaraki prefecture and the TX route. As

¹⁴ Each region is composed of municipalities as follows: the Central (the cities of Kasama, Mito, and Omitama and the towns of Ibaraki, Oarai, and Shirosato), South (the cities of Inashiki, Ishioka, Kasumigaura, Ryugasaki, Thukubamirai, Toride, Tsuchiura, Tsukuba, and Ushiku; the village of Miho; and the towns of Ami and Kawachi), West (the cities of Bando, Chikusei, Joso, Koga, Sakuragawa, Shimotsuma, and Yuki and the town of Yachiyo), North (the cities of Hitachi, Hitachinaka, Hitachiomiya, Hitachiota, Kitaibaraki, Naka, and Takahagi; the town of Daigo; and the village of Tokai), and Rokko regions (the cities of Hokota, Itako, Kamisu, Kashima, and Namekata).

there is a difference in proximity to the new commuting train among the municipalities, the impact of TX's opening might only affect municipalities greater access to the new railway.

We consider the possibility that the city of Tsukuba containing the Tsukuba terminal, the terminus of TX, and its surrounding areas are independent from the other area of Ibaraki prefecture. Kanemoto and Tokuoka (2002) define a Japanese core metropolitan area as an area including a core city with a Density Inhabited District (DID) and with more than 10 thousand residents and with suburban municipalities with more than 10% of workers commuting to the core city. Ibaraki prefecture includes three core metropolitan areas (the Mito, Hitachi and Tsukuba metropolitan areas) according to the National Census (2010). Although each municipality can be contained in multiple metropolitan areas according to the definition provided by Kanemoto and Tokuoka, no municipalities overlap across the Mito, Hitachi and Tsukuba metropolitan areas. Given this we assume that the metropolitan areas are strongly independent from one another. Since TX passes the Tsukuba metropolitan area but does not pass the Mito and Hitachi metropolitan areas, the establishment of TX would especially affect municipalities in the Tsukuba metropolitan area.

Japan experienced a period involving numerous mergers of municipalities known as the great merger of Heisei. A peak occurred in 2005; on March 31th, 2004 there were 3,132 municipalities in Japan, and the number had decreased to 1,821 by March 31th, 2006. Also in Ibaraki, the number of municipalities decreased from 85 to 44 in the period we focus on in this study. To control for these municipal mergers, we add each variable for municipalities to variables for municipalities involved in a merger for the period before the merger.

To observe the effects of proximity to TX on the magnitude of the impact of TX opening, we introduce a cross-term of TimeDistance_{it} and an indicator that identifies whether a municipality is positioned close to TX. The estimation model is as follows:

$$\text{Pov}_{it} = \alpha_{it} + \beta \text{TimeDistance}_{it} + \gamma \text{TimeDistance}_{it} \times \text{ClosetoTX}_i + \delta \mathbf{X}_{it} + \eta_i + \varepsilon_{it}. \quad (1)$$

0)

We define regions closer to TX as the West and South regions; when a municipality is located in one of the regions, the indicator becomes one.

Table 2.4 shows summary statistics for our panel analysis. We find from the table that time distance to the nearest urban metropolitan area of a municipality close to TX dramatically decreased after TX establishment. On the other hand, in municipalities not positioned close to TX, time distance to the nearest urban metropolitan area of a municipality slightly decreased. We consider the possibility that

decreases in accessibility to urban metropolitan areas may improve poverty conditions in municipalities close to TX.

Table 4. Summary Statistics of the Panel Analysis

	Municipalities Not Close to TX				Municipalities Close to TX			
	mean	sd	min	max	mean	sd	min	max
<i>Before Opening TX (year<=2005)</i>								
<i>DISTANCE to the Nearest Urban Metropolitan Area</i>								
Time Distance by Public Transport	73.0	27.3	28	151	42.3	17.7	10	68
<i>POVERTY</i>								
Poverty Rates (%)	0.43	0.17	0.13	0.98	0.29	0.10	0.09	0.46
<i>Other Regional Characteristics</i>								
Population	61,226	53,748	9,873	262,603	89,874	51,627	18,024	200,528
Share of Population Under 15	14.6	1.3	10.8	17.2	14.3	1.8	9.9	17.2
Share of Population Over 65	20.1	4.1	12.6	34.2	16.0	3.2	10.4	20.9
Share of High School Graduates	37.7	2.5	31.3	42.3	35.0	4.2	24.9	41.2
Share of University Graduates	6.7	2.5	3.2	13.9	13.3	4.3	5.7	20.3
Share of Workers in Primary Sector	5.5	3.5	0.7	15.7	3.9	1.7	1.4	7.2
Share of Workers in Industry Sector	30.4	6.3	16.5	45.1	30.0	5.8	16.4	39.1
	N=30	Municipality=15			N=58	Municipality=29		
<i>After Opening TX (year>=2010)</i>								
<i>DISTANCE to the Nearest Urban Metropolitan Area</i>								
Time Distance by Public Transport	69.2	25.0	26	139	27.1	18.5	5	67
<i>POVERTY</i>								
Poverty Rates (%)	0.69	0.35	0.18	2.07	0.50	0.22	0.17	0.95
<i>Other Regional Characteristics</i>								
Population	59,193	54,301	8,786	270,783	93,081	55,426	16,313	226,963
Share of Population Under 15	12.3	1.6	8.5	16.8	13.0	1.8	8.7	16.2
Share of Population Over 65	26.5	4.2	16.7	40.4	23.8	5.5	14.3	38.2
Share of High School Graduates	40.8	4.7	29.2	50.7	35.2	6.9	20.1	47.2
Share of University Graduates	8.7	2.8	4.5	16.1	16.5	5.1	7.5	26.6
Share of Workers in Primary Sector	11.1	7.0	2.9	31.9	11.2	6.5	3.5	30.7
Share of Workers in Industry Sector	30.2	6.0	19.4	45.9	28.7	4.4	19.2	35.9
	N=30	Municipality=15			N=58	Municipality=29		

5.3. The results

Table 2.5 presents the results of our panel analysis. Columns (1)–(2) show the results of the analysis for the whole sample (2000, 2005, 2010, and 2015). In column (1), the coefficient of TimeDistance_{it} is positive and insignificant. This shows that time distance to the nearest urban metropolitan area does not affect rural poverty rates. We then include the interaction term between TimeDistance_{it} and CloseToTX_i in our model as well as TimeDistance_{it} . In column (2), the coefficient of TimeDistance_{it}

is negative and insignificant, and the coefficient of the cross-term is positive with 10% significance. The magnitude of the interaction term denotes that one-minute decreases of TimeDistance_{it} reduce rural poverty rates by roughly 0.006 percentage points.

From columns (1) and (2) we find that accessibility to the nearest urban metropolitan area does not have significant impacts on rural poverty rates overall in the Ibaraki prefecture. However, column (2) suggests that time distance to urban metropolitan areas affects rural poverty in municipalities that are closer to TX. From these results we find that the opening of TX affected only those municipalities with greater accessibility to TX. Although time distance to the nearest urban area also changed in regions not close to TX, their poverty rates do not reflect changes in accessibility. We consider the possibility that a slight change in accessibility to urban metropolitan areas cannot spread or strengthen agglomeration spillover effects enough to improve rural poverty conditions while the opening of TX caused a significant improvement in accessibility to urban metropolitan areas, which was sufficient to improve poverty rates in areas peripheral to TX.

Table 5. Results of the Panel Analysis

	(1)	(2)	(3)	(4)	(5)	(6)
	<i>Baseline Sample</i>		<i>2005 and 2010</i>		<i>2005 and 2015</i>	
Time Distance to the Nearest Urban Metropolitan Area	0.0015 (0.0017)	-0.0020 (0.0026)	-0.00023 (0.0013)	0.00014 (0.0029)	0.0025 (0.0035)	-0.0056 (0.0034)
Time Distance to the Nearest Urban Metropolitan Area × Municipality Close to TX	–	0.0057* (0.0030)	–	-0.00045 (0.0032)	–	0.012** (0.0052)
<i>X_i</i>	Y	Y	Y	Y	Y	Y
Municipality Fixed Effect	Y	Y	Y	Y	Y	Y
Year Fixed Effect	Y	Y	Y	Y	Y	Y
Constant	7.56 (6.78)	5.34 (6.88)	6.52 (6.61)	6.24 (7.20)	6.74 (10.37)	2.78 (9.89)
Adjusted R ²	0.70	0.71	0.41	0.41	0.67	0.70
Municipality	44	44	44	44	44	44
N	176	176	88	88	88	88

Signif. codes: 0 ‘***’ 0.01 ‘**’ 0.05 ‘*’ 0.1. We adopt a cluster robust standard error.

Cluster robust standard errors are in parentheses.

From Table 2.4 we find that while municipalities close to TX decrease their time distance to the nearest urban metropolitan area by roughly 15.2 minutes on average (from 42.3 to 27.1), municipalities far from TX decrease their time distance by only approximately 3.8 minutes (from 73.0 to 69.2). This difference in magnitude may reflect the municipalities’ accessibility to TX, and the poverty conditions of municipalities close to TX might be more sensitive to the opening of TX than those of municipalities

positioned far from TX. Roberto (2008) empirically investigates the case of the United States and finds that the working poor are subjected to a greater burden of commuting costs than the national median in eight metropolitan cities. Their result suggests that those living in poverty experience less price elasticity in commuting costs than those not living in poverty. When improved commuting costs to the nearest urban metropolitan area are still too expensive for those living in poverty, municipalities' poverty conditions may not reflect changing levels of accessibility to urban areas. From this discussion we consider the possibility that a slight change in accessibility to urban areas cannot affect rural poverty rates significantly.

From columns (4) and (6) of Table 2.5 we find a time lag between the opening of TX and changes in regional poverty levels. Chandra and Thompson (2000) focus on the lag of firm relocation after improvement are made in interstate transportation in the United States. They find that impacts of new interstate highways on suburban regions' labor demands do not appear to be significant until several years later. Their results suggest that improvements in accessibility to a closely located urban metropolitan area spread agglomeration spillover effects and increase labor demand in surrounding regions but that lags occur before the effects appear. Provided that lags occur in cases of opening commuting trains, an improvement in poverty rates may not be observed until several years after the opening of TX in the surrounding areas.

To observe the lags of effects appearing after TX opening, we conduct panel analyses (3)–(6). Columns (3)–(4) show the results of analyses conducted on our sample for 2005 and 2010. On the other hand, columns (5)–(6) show the results of analyses conducted on our sample for 2005 and 2015. Columns (3) and (5) adopt time distance to the nearest urban metropolitan area as a distance variable. Columns (4) and (6) further apply the cross-term between ClosetoTX_i and TimeDistance_{it} , as well as TimeDistance_{it} .

In column (3) the coefficient of TimeDistance_{it} is negative and insignificant. In column (4) the coefficient of TimeDistance_{it} is positive and insignificant, and one of the cross-terms is negative and insignificant. In column (5) the coefficient of TimeDistance_{it} is positive and insignificant. In column (6) the coefficient of TimeDistance_{it} is negative and insignificant, and one of the cross-terms is positive with 5% statistical significance.

Columns (3)–(4) show that time distance to urban metropolitan areas does not have significant effects on rural poverty rates in 2005 and 2010. However, from columns (5) and (6) we find that one-minute increases of time distance to the nearest urban metropolitan area cause rural poverty rates to increase by approximately 0.012 percentage points in municipalities close to TX. These results show that the economic effects of TX establishment on peripheral regions did not appear until 6–10 years later. Our

results are consistent with the findings of Chandra and Thompson (2000), which suggest that the effects of improvements in accessibility to urban metropolitan areas take some years to be observed.

However, as another explanation of the results of our panel analyses, TX opening may have spurred inflows of wealthy people to municipalities proximal to TX through an effect called gentrification. If the opening of TX attracted those not living in poverty to the Tsukuba metropolitan area, it may have decreased regional poverty rates even if the number of the people living in poverty did not change. To determine whether the opening of TX decreased the number of low income residents living in areas proximal to TX, we conduct panel analyses focusing on the impacts of TX establishment on the number of households of each income level. If improvements in accessibility to the nearest urban metropolitan area with TX establishment increased regional labor demand in municipalities proximal to TX, the number of low income households should have decreased in municipalities after the opening of TX, decreasing rural poverty rates.

Table 2.6 shows summary statistics for the number of households classified by annual income: less than 3,000,000 yen, 3,000,000 - 5,000,000 yen, 5,000,000 - 10,000,000 yen and more than 10,000,000 yen.¹⁵ Data are also divided between periods before and after TX opening, by the locations of municipalities, and by proximity to TX. From the table we find that the number of households earning less than 3,000,000 yen, 3,000,000 - 5,000,000 yen, and 5,000,000 - 10,000,000 yen increased after the opening of TX for both groups of municipalities. On the other hand, the number of households earning more than 10,000,000 yen decreased after TX establishment in both groups.

¹⁵ We obtain municipal-level data for households divided by income level from The House and Land Statistics Survey conducted by the Ministry of Internal Affairs and Communications Statistics Bureau.

Table 6. Distribution of Household Income in Ibaraki Prefecture

	Municipalities Not Close to TX				Municipalities Close to TX			
	mean	sd	min	max	mean	sd	min	max
<i>Before Opening TX (year<=2005)</i>								
The number of Households Earning less than 3,000,000 yen	5,506.0	7,072.1	0	31,060	5,200.7	4,727.7	0	23,590
The number of Households Earning between 3,000,000-5,000,000 yen	5,156.0	6,211.4	0	24,940	5,269.5	4,236.3	0	18,150
The number of Households Earning between 5,000,000-10,000,000 yen	6,332.4	7,391.2	0	28,500	7,185.2	5,887.8	0	22,670
The number of Households Earning more than 10,000,000 yen	1,792.4	2,347.1	0	10,120	2,356.9	2,160.9	0	9,320
	N=30	Municipality=15			N=58	Municipality=29		
<i>After Opening TX (year>=2010)</i>								
The number of Households Earning less than 3,000,000 yen	8,456.9	8,766.6	2,280	41,140	7,615.7	5,876.5	1,360	26,210
The number of Households Earning between 3,000,000-5,000,000 yen	6,886.7	6,907.2	1,510	30,650	6,793.6	4,902.2	1,380	22,070
The number of Households Earning between 5,000,000-10,000,000 yen	7,450.5	7,316.7	1,330	31,210	7,861.4	6,525.3	1,560	31,010
The number of Households Earning more than 10,000,000 yen	1,618.8	1,605.4	240	6,780	1,736.7	1,486.5	320	6,820
	N=30	Municipality=15			N=58	Municipality=29		

We conduct a panel analysis to investigate the impacts of TX establishment on the number of households of each income level. The estimation models are as follows:

$$\ln \text{Households}_{it}$$

$$= \alpha_{it}$$

$$+ \beta \text{TimeDistance}_{it}$$

$$+ \gamma \text{TimeDistance}_{it} \times$$

$$\text{ClosetoTX}_i + \eta_i + \varepsilon_{it}. \quad (11)$$

In estimation model (11), $\ln \text{Households}_{it}$ is the natural logarithmic of the number of municipality i 's households classified by income level in year t .

Table 2.7 presents the results of the panel analyses based on estimation model (11). Column (1) shows the results of the analysis adopting the natural logarithm of the number of households earning less than 3,000,000 yen as the explained variable. In column (1), the coefficient of time distance to the nearest urban metropolitan area is negative with one percent significance. The magnitude shows that one-minute increases in TimeDistance_{it} reduce the number of households earning less than 3,000,000 yen by roughly 0.007 percentage points. Then, the coefficient of the cross-term is positive with five percent

significance. The magnitude of the interaction term denotes that one-minute decreases of TimeDistance_{it} reduce the number of households earning less than 3,000,000 yen by roughly 0.008 percentage points in municipalities proximal to TX.

Column (2) shows the results of our estimation adopting the natural logarithm of the number of households earning between 3,000,000 and 5,000,000 yen as a dependent variable. In column (2) the coefficient of TimeDistance_{it} is negative with five percent significance. The magnitude denotes that one-minute increases in TimeDistance_{it} reduce the number of households earning between 3,000,000 and 5,000,000 yen by roughly 0.009 percentage points. The coefficient of the cross-term is positive with 10% significance. The magnitude of the interaction term denotes that one-minute decreases of TimeDistance_{it} reduce the number of households earning between 3,000,000 and 5,000,000 yen by roughly 0.007 percentage points in municipalities proximal to TX.

Column (3) uses the natural logarithm of the number of households earning between 5,000,000 and 10,000,000 yen as the explained variable and shows similar results to those of columns (1) and (2) for points of coefficients of TimeDistance_{it} and the cross-term. The coefficient of time distance to the nearest urban metropolitan area is negative with one percent significance. The magnitude denotes that one-minute increases in TimeDistance_{it} reduce the number of households earning between 3,000,000 and 5,000,000 yen by roughly 0.009 percentage points. Then, the coefficient of the cross-term is positive with 10% significance. The magnitude of the interaction term denotes that one-minute increases of TimeDistance_{it} increase the number of households earning between 5,000,000 and 10,000,000 yen by roughly 0.005 percentage points in municipalities proximal to TX.

Column (4) describes the results of an analysis adopting the natural logarithm of the number of households earning more than 10,000,000 yen as a dependent variable. In column (4), the coefficient of TimeDistance_{it} is negative and insignificant. Then, the coefficient of the cross-term is positive and insignificant. These results show that time distance to the nearest urban metropolitan area does not affect the number of households earning more than 10,000,000 yen.

Regarding TimeDistance_{it} , the results of columns (1), (2) and (3) show that increases in time distance to the nearest urban metropolitan area decrease the number of households earning less than 3,000,000 yen, 3,000,000 to 5,000,000 yen, and 5,000,000 to 10,000,000 yen across Ibaraki prefecture. From these results we find that households with incomes of less than 10,000,000 yen tend to relocate to areas with better accessibility to the nearest urban metropolitan area to lower commuting costs. Moreover, we also find that higher income households sensitively react to lower time distances to the nearest urban metropolitan area. This is consistent with our above discussion showing that lower income households enjoy less mobility than higher income households. From the above discussion, we note that

improvements in accessibility to the nearest urban metropolitan areas attract households earning less than 10,000,000 yen, and the scale of impacts depends on household income levels. However, column (4) suggests that the number of households earning more than 10,000,000 yen (the wealthiest cohort included in our sample) was not affected by time distance to the nearest urban metropolitan area. We posit that since households with incomes of more than 10,000,000 yen have enough income to reside in desirable areas, they do not respond to improvements in accessibility to the nearest urban metropolitan area with relocation.

Regarding the cross-term, columns (1) – (3) show that increases in time distance to the nearest urban metropolitan area increase the number of households with incomes of less than 3,000,000 yen, of 3,000,000 and 5,000,000 yen, and of 5,000,000 and 10,000,000 yen for municipalities proximal to TX. These results are consistent with our assumption that improvements in accessibility to the nearest urban metropolitan area increase regional income levels and reduce the number of lower income households earning less than 3,000,000 yen in the sample. The results also show that the number of households earning between 3,000,000 and 5,000,000 yen and between 5,000,000 and 10,000,000 yen decrease with improvements in accessibility to the nearest urban metropolitan area. Since there is a tendency for lower income households to be strongly affected by time distance to the nearest urban metropolitan area, we except the reduction in the number of households earning between 3,000,000 and 5,000,000 yen and between 5,000,000 and 10,000,000 to be partly canceled out by lower income households before TX establishment and for incomes to increase after TX establishment.

The above results show that the opening of TX decreased the number of low income households in municipalities proximal to TX. We also observe that improvements in accessibility to the nearest urban metropolitan area attract households earning less than 10,000,000 yen. These results suggest that improvements in accessibility to the nearest urban metropolitan area decrease regional poverty levels by increasing regional income levels while activating economic activity in these areas through inflows of households from other areas.

Table 7. Results of Opening TX on the Number of Households Divided by their Income

	(1)	(2)	(3)	(4)
	ln (Households Earning less than 3,000,000 yen)	ln (Households Earning between 3,000,000- 5,000,000 yen)	ln (Households Earning between 5,000,000- 10,000,000 yen)	ln (Households Earning more than 10,000,000 yen)
Time Distance to the Nearest Urban Metropolitan Area	-0.00727*** (-3.690)	-0.00853** (-2.677)	-0.00885*** (-4.018)	-0.00278 (-1.375)
Time Distance to the Nearest Urban Metropolitan Area × Municipalities Close to TX	0.00782** (2.581)	0.00738* (1.993)	0.00486* (1.686)	0.00355 (1.603)
Municipality Fixed Effects	Y	Y	Y	Y
Year Fixed Effect	Y	Y	Y	Y
Constant	8.287*** (198.8)	8.416*** (152.1)	8.826*** (175.1)	7.486*** (86.76)
Observations	157	157	157	157
R2	0.780	0.563	0.313	0.340
N	42	42	42	42

Robust t-statistics in parentheses

*** p<0.01, ** p<0.05, * p<0.1

However, it may be that decreases in the number of lower income households are a mere result of those living in poverty being driven out by increased housing rents resulting from gentrification. If gentrification caused increases in housing rents in areas proximal to TX, rendering rents too expensive for the poor, those in poverty would have needed to relocate to other areas with lower housing rents, and we should observe decreases in the number of lower income households in the Tsukuba metropolitan area.

We believe that those living in poverty before TX opening were not driven from their residences even after the opening of TX for two main reasons. First, as a legal condition protecting those in poverty from being driven out, the Leased Land and House Lease Law is enforced in Japan. The law provides that owners can increase rents only when "*the building rent becomes unreasonable, as a result of the increase or decrease in tax and other burdens relating to the land or the buildings, as a result of the rise or fall of land or building prices or fluctuations in other economic circumstances, or in comparison to the rents on similar buildings in the vicinity, the parties may, notwithstanding the contract conditions*".¹⁶ Moreover, the law allows residents to negotiate with the owners through civil conciliation and to deposit their rents in the same amount before rent levels change. Therefore, we conclude that those living in poverty were not immediately driven out even after gentrification occurred.

Second, since relocation costs place more serious burdens on those living in poverty than on those not in poverty, the poor tend to remain in their residences for longer periods of time. Table 2.8 shows the number and share of households classified by monthly housing rent levels and periods of residence in

¹⁶ Article 32 (1) of the Leased Land and House Lease Law.

Ibaraki prefecture in 2013.¹⁷ We find that for those with housing rents of 0 - 10,000 yen and 10,000 – 20,000 yen, more than 30% of households had resided in their current residences for 13 or more years. For households with housing rents of 20,000 – 40,000 yen, the share of those living in the same residence for more than 13 years is roughly 24.6%, and this share decreases with increases in housing rent. Since housing rent is very positively related to household income, we can conclude that households living in poverty enjoy less mobility than those not in poverty.

Table 8. The Number and Share of Households Divided by Monthly Rent and Duration of Residence in 2013

	All	0 yen	0-10,000 yen	10,000- 20,000 yen	20,000- 40,000 yen	40,000- 60,000 yen	60,000- 80,000 yen	80,000- 100,000 yen	100,000- 150,000 yen	150,000- 200,000 yen
All	209,400 (100)	4,900 (100)	10,900 (100)	13,800 (100)	46,800 (100)	78,900 (100)	43,500 (100)	7,500 (100)	2,900 (100)	200 (100)
0-2 years	86,100 (41.1)	1,900 (38.8)	4,100 (37.6)	5,200 (37.7)	17,300 (37)	33,300 (42.2)	19,900 (45.7)	3,100 (41.3)	1,200 (41.4)	100 (50)
3-7 years	63,700 (30.4)	1,100 (22.4)	2,000 (18.3)	2,800 (20.3)	12,600 (26.9)	25,600 (32.4)	15,500 (35.6)	3,200 (42.7)	800 (27.6)	100 (50)
8-12 years	21,800 (10.4)	400 (8.2)	1,100 (10.1)	1,400 (10.1)	5,300 (11.3)	8,500 (10.8)	4,000 (9.2)	700 (9.3)	400 (13.8)	0 (0)
13-22 years	21,500 (10.3)	600 (12.2)	1,700 (15.6)	1,700 (12.3)	5,700 (12.2)	7,700 (9.8)	3,400 (7.8)	400 (5.3)	300 (10.3)	0 (0)
more than 23 years	16,200 (7.9)	1,000 (20.4)	1,900 (17.4)	2,700 (19.6)	5,800 (12.4)	3,800 (4.8)	700 (1.6)	200 (2.7)	100 (3.4)	0 (0)

Share of the period that households live in their house divided by monthly rent in parentheses

From the above discussions we conclude that many households living in poverty before TX establishment continued to live in the same rental units even after TX opening, and the number of lower income households decreased in areas proximal to TX after TX opening. Therefore, we conclude that improvements in accessibility to the nearest urban metropolitan area improve the living standards of those living in poverty in surrounding areas, and our results suggest that decreases in poverty rates resulting from TX establishment (described in Table 2.5) are not a mere result of increases in the number of residents not living in poverty due to gentrification.

In addition, we discuss changes of trends of commuting activities in Ibaraki Prefecture after TX opening. Figure 2.3 shows transition of the share of commuters to municipalities where they live in or to Tokyo.¹⁸ In the graphs, municipalities in Ibaraki Prefecture are classified by proximity to TX. The left graph suggests that, in municipalities not closely located to TX, the share of commuters to the municipalities where they live in decreased more than in municipalities close to TX after opening TX. From the right one, we can observe that the share of commuters to Tokyo more greatly increased in municipalities not close to TX than in municipalities close to TX after TX opening. From those data, we consider that TX opening did not increase commuters to Tokyo in municipalities close to TX compared to

¹⁷ Data are drawn from The House and Land Statistics Survey conducted by the Ministry of Internal Affairs and Communications Statistics Bureau (2018).

¹⁸ We refer the Japanese national Census (2000, 2005, 2010 and 2015)

municipalities not close to TX. The decreases in the poor in municipalities close to TX might be caused not by better commuting condition to the metropolitan areas, but by expand of ranges of agglomeration spillover effects and increases in regional labor demand.

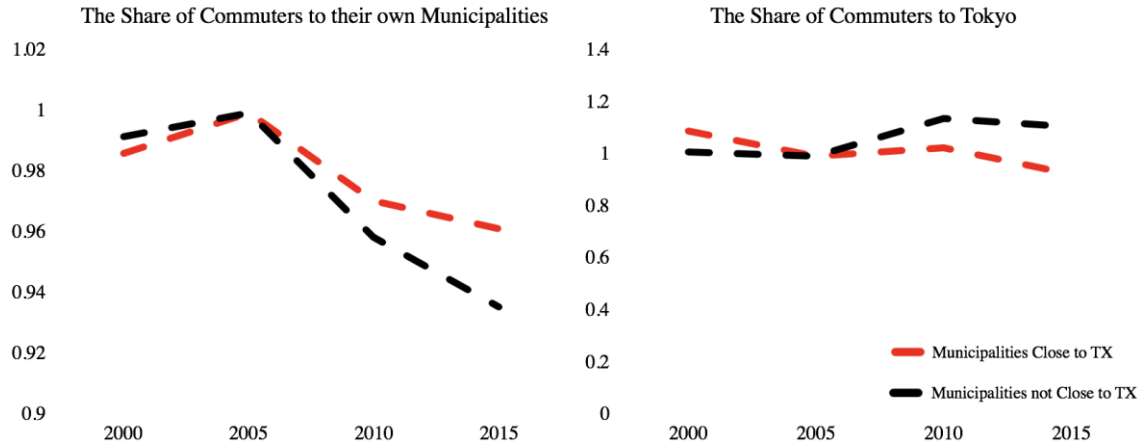


Figure 3. Transitions of the Share of Commuters Classified by Commuting Places

6. Conclusion

We investigate the relationship between municipalities' accessibility to the nearest urban metropolitan area and rural poverty rates by focusing on the case of Japanese municipalities. We find that while increasing the time distance to the nearest urban metropolitan area increases rural poverty rates, linear distance does not explain rural poverty. When focusing on a case involving many geographical barriers, there is a difference between linear distance and time distance in explaining the spillover effects of urban agglomeration on regional economic performance.

Moreover, we focus on the impacts of the opening of commuting train TX in 2005 on surrounding municipalities' poverty rates and we conduct a panel analysis to control for time invariant characteristics of municipalities. Even when we control for regional time invariant characteristics, the causal effects of access to the nearest urban metropolitan area on rural poverty rates are significant for municipalities positioned close to the new commuting train. This result suggests that slight changes in accessibility to the nearest urban metropolitan area do not significantly affect regional poverty levels.

From our estimations, one-minute increases in time distance to the nearest urban metropolitan area increase rural poverty rates by roughly 0.002 percentage points. This implies that roughly 0.78 additional households begin to receive public assistance, increasing annual expenditures made on public assistance by roughly 1.75 million yen on average. In addition, we find that the economic effects of TX establishment on surrounding regions did not appear until 6–10 years later. This result suggests that when

governments invest in transportation infrastructures to improve their economic performance, they ought to expect effects to appear only a few years later.

Our results show that economic agglomeration spillovers are effective in reducing poverty levels in surrounding regions and that improved accessibility to proximal urban metropolitan areas increases the magnitude and range of effects. Transportation investments that improve levels of accessibility to urban metropolitan areas may stimulate economic performance in such areas and reduce poverty levels.

REFERENCES

- Audretsch, D. B., E. E. Lehmann, and S. Warning (2005) "University Spillovers and New Firm Location", *Research Policy*, Vol. 34, No. 7, pp. 1113-1122.
- Boscoe, F. P., K. A. Henry, and M. S. Zdeb (2012) "A Nationwide Comparison of Driving Distance Versus Straight-Line Distance to Hospitals", *The Professional Geographer*, Vol. 64, No. 2, pp. 188-196.
- Chandra, A. and E. Thompson (2000) "Does Public Infrastructure Affect Economic Activity?: Evidence from the Rural Interstate Highway System", *Regional Science and Urban Economics*, Vol. 30, No. 4, pp. 457-490.
- Chandy, L., and C. Smith (2014). "How poor are America's poorest? US \$2 a day poverty in a global context". *Global Economy and Development, Global Views, Policy Paper*, 3.
- Clark, W. A., Y. Huang and S. Withers (2003). "Does commuting distance matter?: Commuting tolerance and residential change". *Regional Science and Urban Economics*, Vol. 33, No. 2, pp. 199-221.
- Deaton, A. (2004) "Measuring Poverty", in A. Banerjee, R. Benabou, and D. Mookherjee, ed., *Understanding Poverty*, Oxford: Oxford University Press, pp. 3-16.
- Förster, M. and M. M. d'Ercole (2005) "Income Distribution and Poverty in OECD Countries in the Second Half of the 1990s", *OECD Social Employment and Migration Working Papers* 2005;22.
- International Labour Organization [ILO] (2016). *World Employment Social Outlook: Trends* 2016. Geneva: International Labour Office.
- Glaeser, E. L., M. E. Kahn, and J. Rappaport (2008). "Why do the poor live in cities? The role of public transportation". *Journal of urban Economics*, Vol. 63, No. 1, pp. 1-24.
- Japan Travel Bureau (2000, 2005, 2010, 2015) *The Timetable Magazine*. JTB Publishing.
- Kanemoto, Y. and K. Tokuoka (2002). "Proposal for the standards of metropolitan areas of Japan". *Journal of Applied Regional Science*, Vol. 7, pp. 1-15.

- Lucas, R. E. (2001) "The Effects of Proximity and Transportation on Developing Country Population Migrations", *Journal of Economic Geography*, Vol. 1, No. 3, pp. 323-339.
- Marshall, A. (1920) "Industry and Trade: A Study of Industrial Technique and Business Organization; and of their Influences on the Conditions of Various Classes and Nations", London: Macmillan.
- Molho, I. (1995) "Migrant Inertia, Accessibility and Local Unemployment", *Economica*, Vol. 62, pp. 123-132.
- Partridge, M. D. and D. S. Rickman (2008) "Distance from Urban Agglomeration Economies and Rural Poverty", *Journal of Regional Science*, Vol. 48, No. 2, pp. 285-310.
- Roberto, E. (2008) "Commuting to Opportunity: The Working Poor and Commuting in the United States", Washington, DC: Brookings Institution.
- Rosenthal S. S. and W. C. Strange (2003) "Geography, Industrial Organization, and Agglomeration", *Review of Economics and Statistics*, Vol. 85, No. 2, pp. 377-393.
- Rosenthal S. S. and W. C. Strange (2004) "Evidence on the Nature and Sources of Agglomeration Economies", in J. V. Henderson and J. F. Thisse, eds., *Handbook of Regional and Urban Economics*, Vol. 4, Amsterdam: Elsevier, pp. 2119-2171.
- Rosenthal S. S. and W. C. Strange (2006) "The Micro-Empirics of Agglomeration Economies", in R. Arnott and D. McMillen, eds., *The Blackwell Companion to Urban Economics*, Oxford: Blackwell, pp 7-23.